

Probabilistic methods for nuclear data

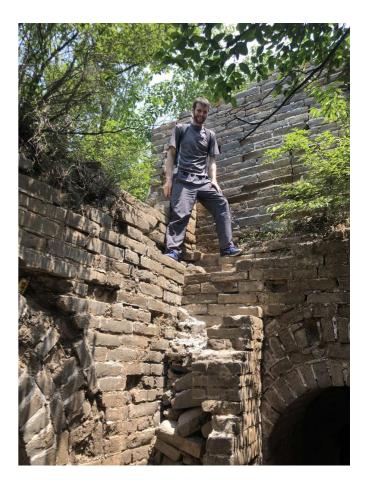
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Workshop on Computational Nuclear Science and Engineering 13 July 2021

Short bio

- Studied physics at TU Vienna
- PhD in nuclear data evaluation 2015
- Postdoc at CEA Saclay (2015-2018) and Uppsala University (2018-2019)
- Since 2020 nuclear physicist in Nuclear Data Section at IAEA dealing with nuclear data library projects and code development



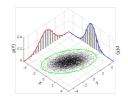
Outline

About nuclear data

Bayesian statistics



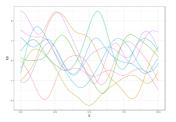
Multivariate normal distribution



Approaches to solve Bayes equation in nuclear data evaluation

Generalized Least Squares

Gaussian processes

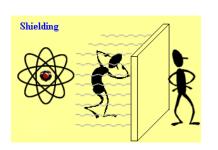


Nuclear data

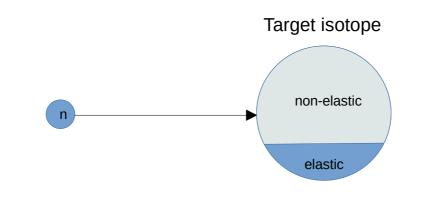


PSI Gantry 2 facility





Probabilities of various nuclear interactions involving the atomic nuclei, e.g., cross sections.



Relevant for:

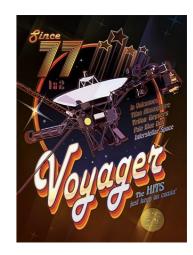
- Reactor physics
- Radiation dosimetry
- Radiation protection
- Radioactive waste management
- Astrophysics
- Nuclear medicine
- Fusion research
- ...



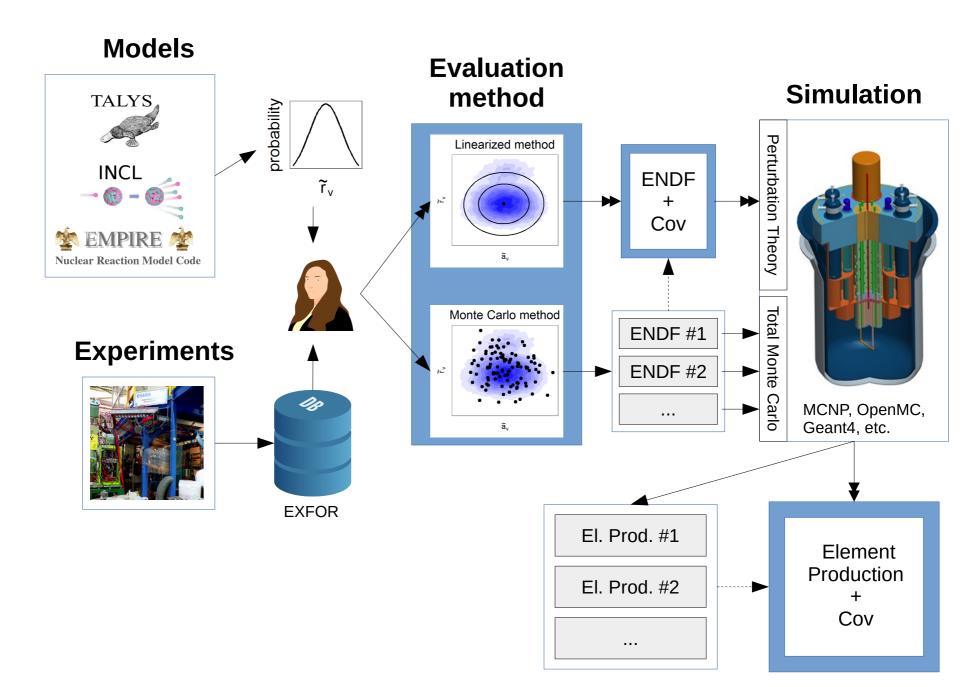
Palisades Nuclear Generating Stations



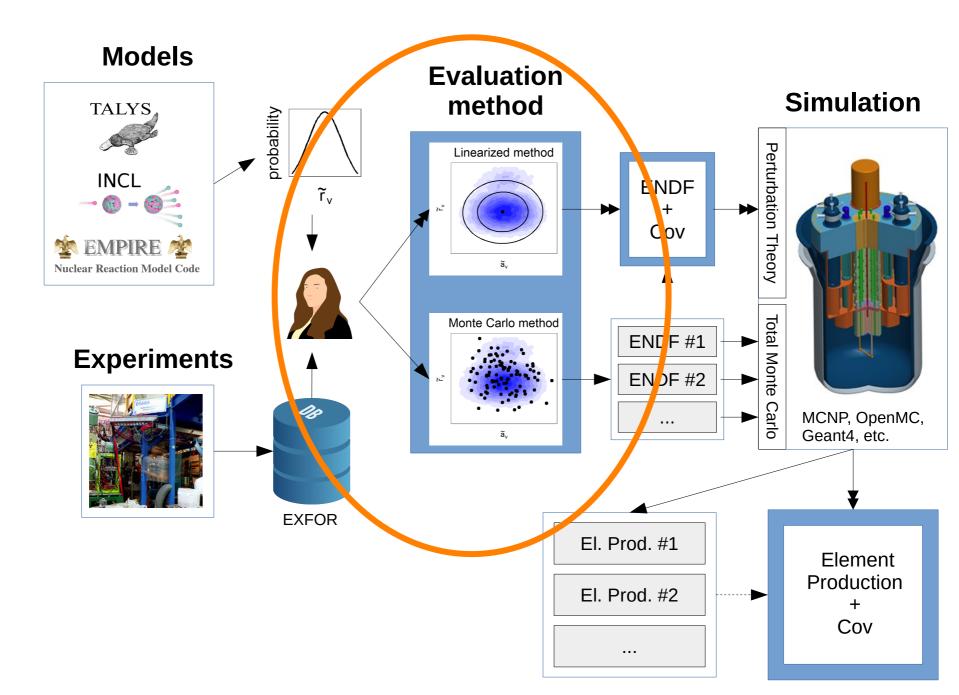
Joint European Torus



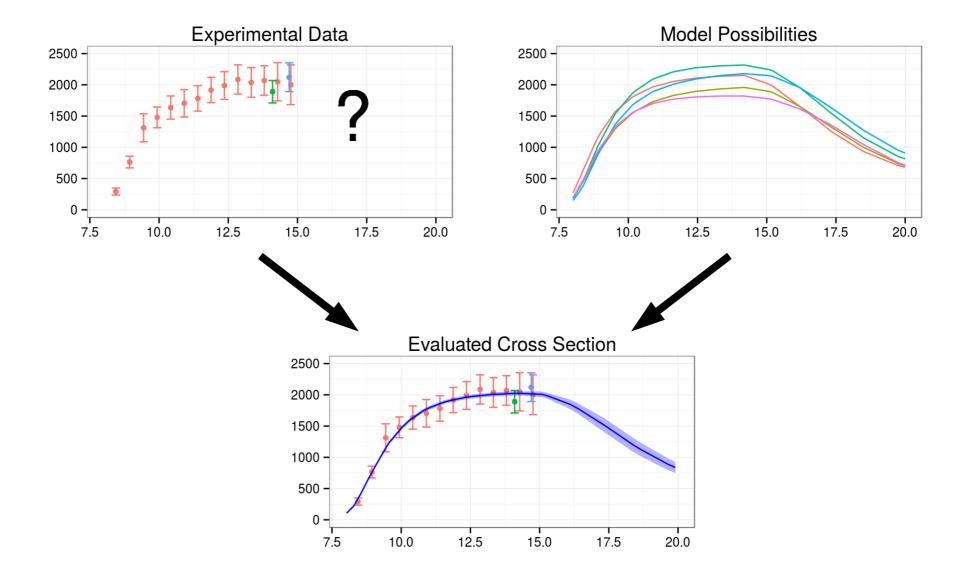
Nuclear data evaluation



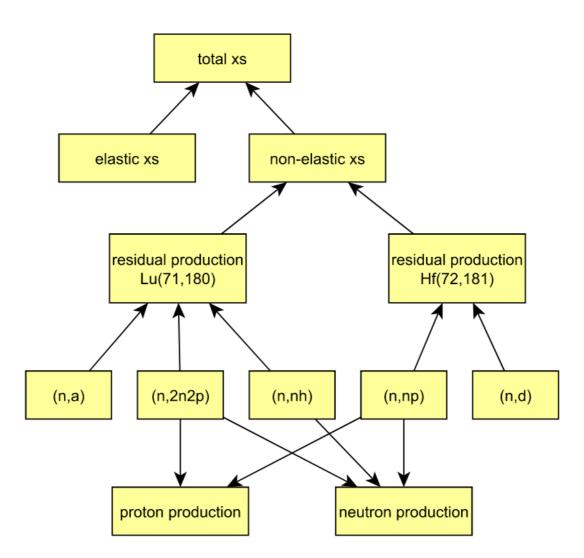
Nuclear data evaluation



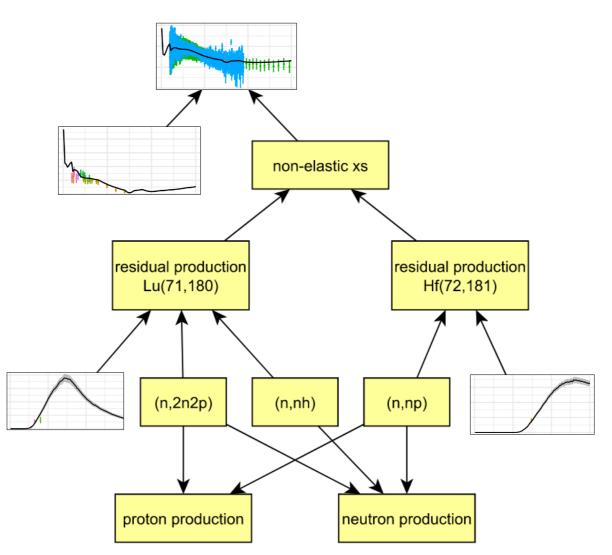
General task setting



System of reactions



System of reactions



Jaynes Robot



Edwin Thompson Jaynes 1922-1998

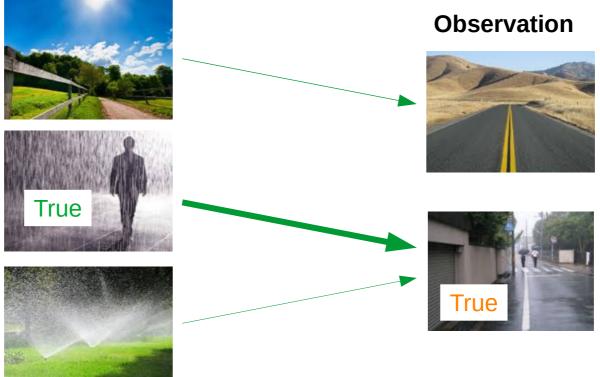
Which hypothesis is true?

Probability Theory The Logic of Science E. T. JAYNES

CAMBRIDGE

Consistency with Aristotelian logic

Hypothesis

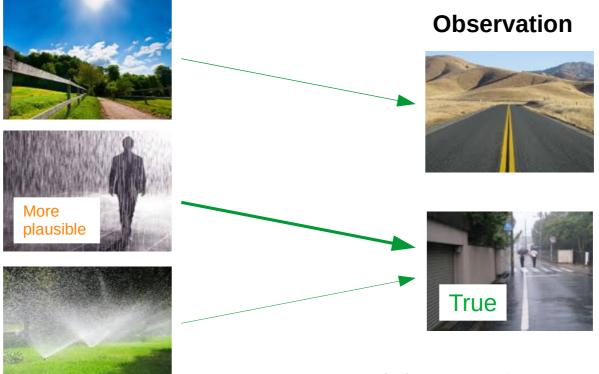


If A is true, then B is true A is true

Therefore, B is true

Consistency with common sense

Hypothesis



If A is true, then B is true B is true

Therefore, A becomes more plausible

Desiderata

(I) Degrees of Plausibility are represented by real numbers

(II) Qualitative Correspondence with common sense

(IIIa) If a conclusion can be reasoned out in more than one way, then every possible way must lead to the same result.

(from E.T. Jaynes, Probability Theory: The Logic of Science)



Scanned at the American Institute of Physics

Richard Threlkeld Cox 1898-1991

Desiderata / Cox theorem

(I) Degrees of Plausibility are represented by real numbers

(II) Qualitative Correspondence with common sense

(IIIa) If a conclusion can be reasoned out in more than one way, then every possible way must lead to the same result.

(from E.T. Jaynes, Probability Theory: The Logic of Science)



Scanned at the American Institute of Physics

Richard Threlkeld Cox 1898-1991

Computation rules of probability theory follow

e.g., product rule

 $P(H, O) = P(O \mid H)P(H)$

Bayesian update formula



Thomas Bayes 1701-1761

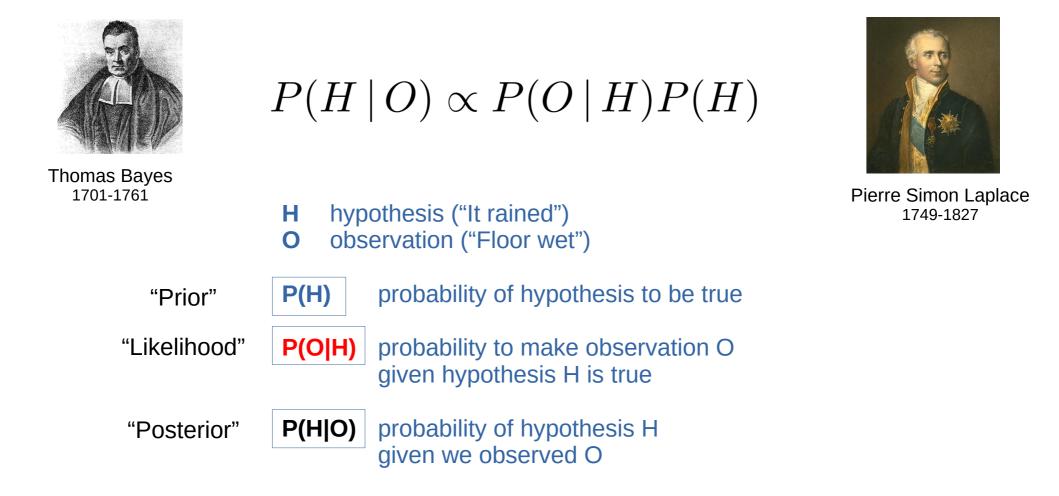
 $P(H \mid O) \propto P(O \mid H)P(H)$

- **H** hypothesis ("It rained")
- **O** observation ("Floor wet")
- **P(H)** probability of hypothesis to be true
- **P(O|H)** probability to make observation O given hypothesis H is true
- **P(H|O)** probability of hypothesis H given we observed O



Pierre Simon Laplace 1749-1827

Bayesian update formula



Bayesian update formula



Thomas Bayes 1701-1761

 $P(H \mid O) \propto P(O \mid H)P(H)$



Pierre Simon Laplace 1749-1827

H hypothesis

O observation

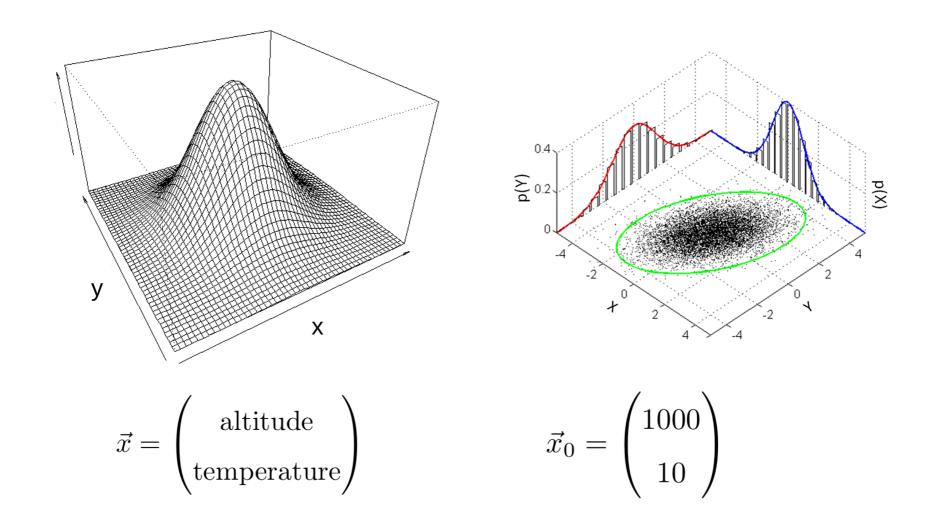
e.g., H(x) := "The cross section at 5 MeV is x mBarn"

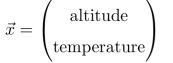
 $P(H(x)) \rightarrow P(x)$

Which probability distributions for P(O|H) and P(H)?

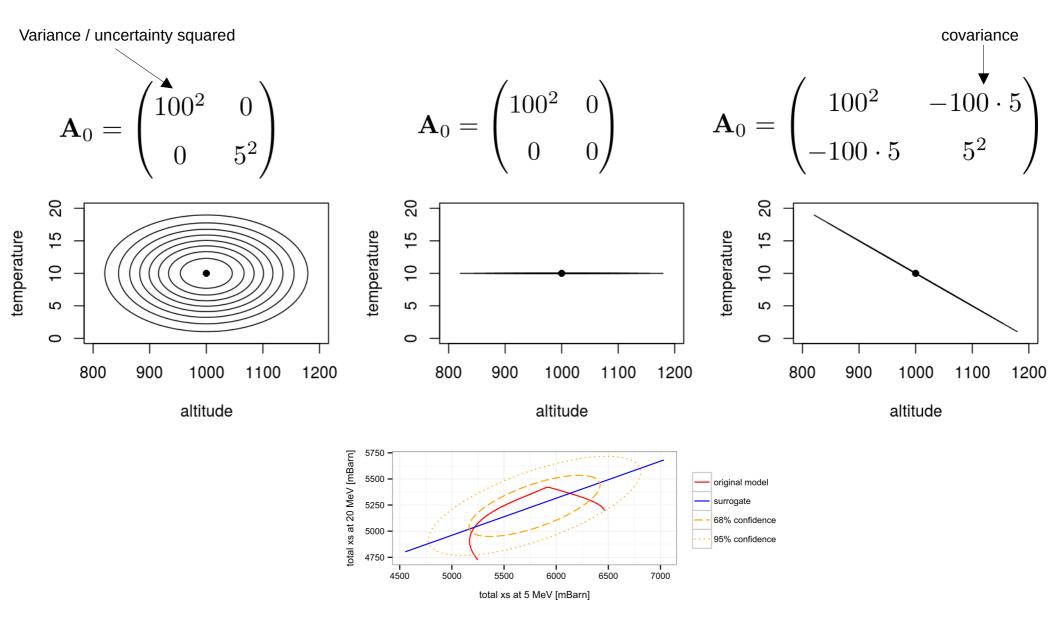
Multivariate normal distribution

$$\mathcal{N}(\vec{x} \,|\, \vec{x}_0, \mathbf{A}_0) = \frac{1}{\sqrt{(2\pi)^N |\mathbf{A}_0|}} \exp\left(-\frac{1}{2} \left(\vec{x} - \vec{x}_0\right)^T \mathbf{A}_0^{-1} \left(\vec{x} - \vec{x}_0\right)\right)$$





Covariance matrix

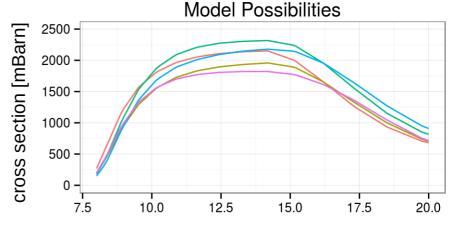


A covariance matrix captures linear relationships and uncertainties

Bayesian Inference in Nuclear Data

Basic assumption

Observables can be predicted by nuclear model M based on parameter set p



incident energy [MeV]

 $\pi(\vec{p} \mid M)$ $f(\vec{\sigma} \mid \vec{p}, M)$ $\pi(\vec{p} \mid \vec{\sigma}, M)$

prior distribution on model parameters

likelihood: probability to measure cross sections $\boldsymbol{\sigma}$ if model parameter set \boldsymbol{p} is true

posterior: refined probability distribution for parameters taking into account observations σ

$\pi(\vec{p} \,|\, \vec{\sigma}, M) \propto f(\vec{\sigma} \,|\, \vec{p}, M) \pi(\vec{p} \,|\, M)$

Which probability distributions?

Prior distribution (model parameters)

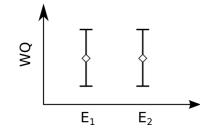
$$\pi(\vec{p} \mid M) = \frac{1}{\sqrt{(2\pi)^d \det \mathbf{A}_0}} \exp\left(-\frac{1}{2}(\vec{p} - \vec{p}_0)^T \mathbf{A}_0^{-1}(\vec{p} - \vec{p}_0)\right)$$

supported by

- principle of maximum entropy
- easy to work with!

Likelihood (experimental data)

$$f(\vec{\sigma} \mid \vec{p}, M) = \frac{1}{\sqrt{(2\pi)^N \det \mathbf{B}}} \exp\left(-\frac{1}{2}(\vec{\sigma} - M[\vec{p}])^T \mathbf{B}^{-1}(\vec{\sigma} - M[\vec{p}])\right)$$



supported by

- principle of maximum entropy,
- limiting distribution,
- central limit theorem
- easy to work with!

Cooking a covariance matrix

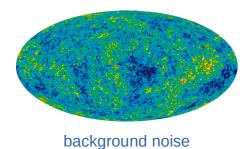
. . .







sample thickness*



"statistical" (counting) uncertainty

$$\mathbf{B}_{\text{stat}} = \begin{bmatrix} \sigma_1^2 \delta_1^2 & 0\\ 0 & \sigma_2^2 \delta_2^2 \end{bmatrix}$$

detector calibration uncertainty

$$\mathbf{B}_{\text{cal}} = \begin{bmatrix} \sigma_1^2 \gamma_1^2 & \sigma_1 \sigma_2 \gamma_1 \gamma_2 \\ \sigma_1 \sigma_2 \gamma_1 \gamma_2 & \sigma_2^2 \gamma_2^2 \end{bmatrix}$$

uncertainty about sample thickness uncertainty about background noise

$$\mathbf{B} = \mathbf{B}_{\text{stat}} + \mathbf{B}_{\text{cal}} + \mathbf{B}_{\text{thick}} + \dots$$

МQ

 E_1

 E_2

Quantities of interest

• Expectation of parameters according to posterior distribution

$$\mathbb{E}\left[\vec{\sigma}\right] = \int \int \cdots \int M(\vec{p}) \pi(\vec{p} \,|\, \vec{\sigma}, M) dp_1 dp_2 \dots dp_N$$

• Mode (maximum) of posterior distribution

find \vec{p}_{max} that maximizes $\pi(\vec{p} \mid \vec{\sigma}, M)$

- Standard deviation parameters according to posterior distribution
- Correlations of distribution

Nuclear data challenges

Regarding nuclear models

- Non-linear (nuclear physics is complicated)
- Not analytic (differential equations, simulation)
- Computationally expensive (minutes to hours)

Regarding linear algebra

• Computing with large covariance matrices

Regarding statistical models

- Imperfect physics model
- Multivariate normal distribution may be not always appropriate
- Uncertainties wrong or unknown

$$\pi(\vec{p} \mid \vec{\sigma}, M) \propto f(\vec{\sigma} \mid \vec{p}, M) \pi(\vec{p} \mid M)$$
$$f(\vec{\sigma} \mid \vec{p}, M) = \frac{1}{\sqrt{(2\pi)^d \det \mathbf{B}}} \exp\left(-\frac{1}{2}(\vec{\sigma} - M[\vec{p}])^T \mathbf{B}^{-1}(\vec{\sigma} - M[\vec{p}])\right)$$

Nuclear data challenges

Regarding nuclear models

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Regarding linear algebra

• Computing with large covariance matrices

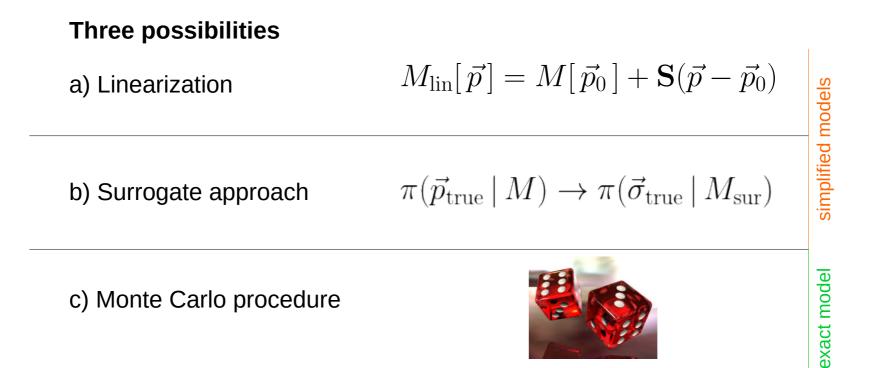
Regarding statistical models

- Imperfect physics model
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$$\pi(\vec{p} \mid \vec{\sigma}, M) \propto f(\vec{\sigma} \mid \vec{p}, M) \pi(\vec{p} \mid M)$$
$$f(\vec{\sigma} \mid \vec{p}, M) = \frac{1}{\sqrt{(2\pi)^d \det \mathbf{B}}} \exp\left(-\frac{1}{2}(\vec{\sigma} - M[\vec{p}])^T \mathbf{B}^{-1}(\vec{\sigma} - M[\vec{p}])\right)$$

How to evaluate?

 $\pi(\vec{p}_{\rm true} \,|\, \vec{\sigma}_{\rm exp}, M) \propto f(\vec{\sigma}_{\rm exp} \,|\, \vec{p}_{\rm true}, M) \pi(\vec{p}_{\rm true} \,|\, M)$



Monte Carlo approach

$$\pi(\vec{p}_{\text{true}} \mid \vec{\sigma}_{\text{exp}}, M_{\text{lin}}) \propto \exp\left[-\frac{1}{2}(\vec{\sigma}_{\text{exp}} - M[\vec{p}_{\text{true}}])^T \mathbf{B}^{-1}(\vec{\sigma}_{\text{exp}} - M[\vec{p}_{\text{true}}])\right] \times \exp\left[-\frac{1}{2}(\vec{p}_{\text{true}} - \vec{p}_0)^T \mathbf{A}_0^{-1}(\vec{p}_{\text{true}} - \vec{p}_0)\right]$$

1) generate ensemble $\vec{p_1}, \vec{p_2}, \vec{p_3}, \dots$ from prior

2) calculate $M[\vec{p_1}], M[\vec{p_2}], M[\vec{p_3}], \dots$



3) Calculate statistics from posterior distribution, e.g., best estimates, correlations, uncertainties, etc. by likelihood weighting $f(\sigma_{exp} | \vec{p_i}, M)$

$$\mathbb{E}[\vec{\sigma}_{\text{true}}] = \frac{\sum_{i=1}^{n} M[\vec{p}_i] f(\sigma_{\text{exp}} | \vec{p}_i, M)}{\sum_{i=1}^{n} f(\sigma_{\text{exp}} | \vec{p}_i, M)}$$

Linearization

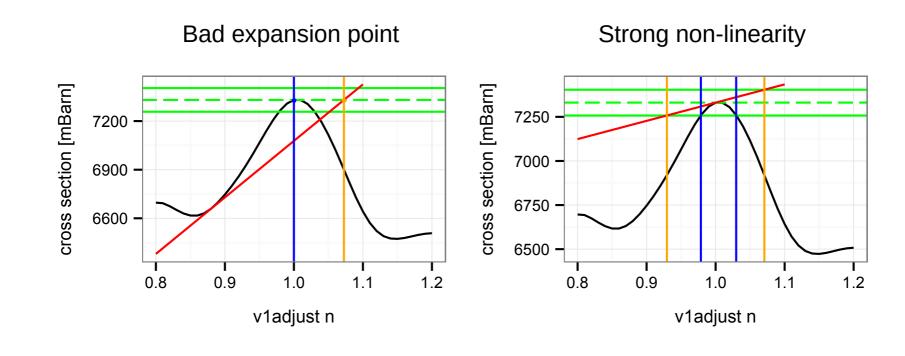
$$\pi(ec{p}_{ ext{true}} \,|\, ec{\sigma}_{ ext{exp}}, M) \propto f(ec{\sigma}_{ ext{exp}} \,|\, ec{p}_{ ext{true}}, M) \pi(ec{p}_{ ext{true}} \,|\, M)$$

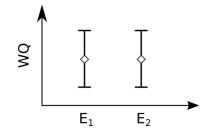
$$\pi(\vec{p}_{\text{true}} \mid \vec{\sigma}_{\text{exp}}, M_{\text{lin}}) \propto \exp\left[-\frac{1}{2}(\vec{\sigma}_{\text{exp}} - M_{\text{lin}}[\vec{p}_{\text{true}}])^T \mathbf{B}^{-1}(\vec{\sigma}_{\text{exp}} - M_{\text{lin}}[\vec{p}_{\text{true}}])\right] \times \exp\left[-\frac{1}{2}(\vec{p}_{\text{true}} - \vec{p}_0)^T \mathbf{A}_0^{-1}(\vec{p}_{\text{true}} - \vec{p}_0)\right]$$

$$M_{\rm lin}[\vec{p}] = M[\vec{p}_0] + \mathbf{S}(\vec{p} - \vec{p}_0)$$

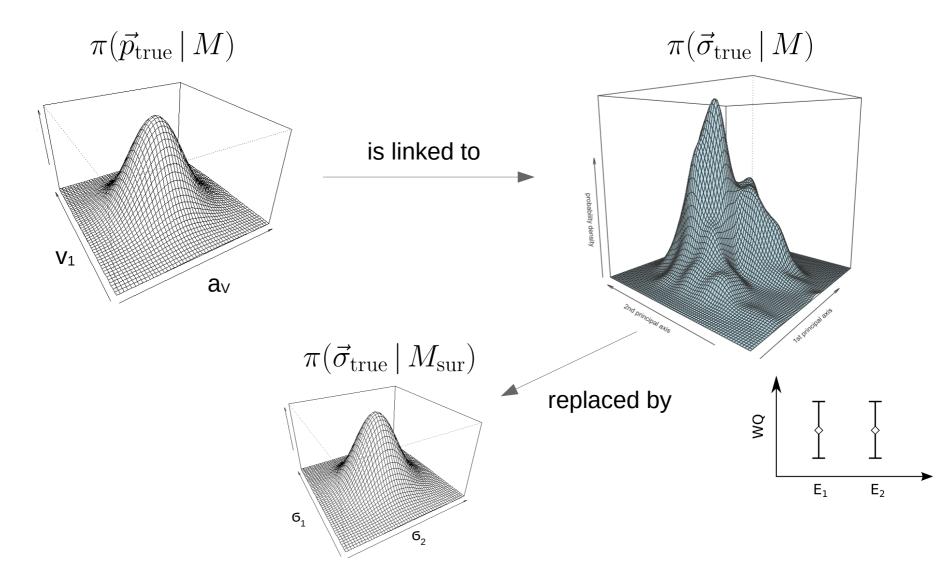
Analytical update formulas	$\mathbf{A}_1 = \mathbf{A}_0 - \mathbf{A}_0 \mathbf{S}^T (\mathbf{S} \mathbf{A}_0 \mathbf{S} + \mathbf{B})^{-1} \mathbf{S} \mathbf{A}_0$
	$\vec{p}_1 = \vec{p}_0 + \mathbf{A}_0 \mathbf{S}^T (\mathbf{S} \mathbf{A}_0 \mathbf{S} + \mathbf{B})^{-1} (\vec{\sigma}_{\exp} - M_{\lim}[\vec{p}_0])$

Linearization: Possible issues





Surrogate approach basic idea



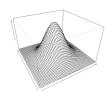
 $\pi(\vec{p}_{\rm true} \,|\, \vec{\sigma}_{\rm exp}, M) \propto f(\vec{\sigma}_{\rm exp} \,|\, \vec{p}_{\rm true}, M) \pi(\vec{p}_{\rm true} \,|\, M)$

Surrogate approach construction of M_{sur}

1) draw ensemble $\vec{p_1} \, \vec{p_2}, \vec{p_3}, \dots, \vec{p_n}$ from prior distribution $\pi(\vec{p_{\text{true}}} \mid M)$ 2) use nuclear model to calculate $\vec{\sigma_1} = M[\vec{p_1}], \vec{\sigma_2} = M[\vec{p_2}], \dots$

3) estimate multivariate normal distribution in observation space

$$\vec{\sigma}_0 = \frac{1}{n} \sum_{i=1}^n \vec{\sigma}_i \qquad \mathbf{A}_0 = \frac{1}{n} \sum_{i=1}^n (\vec{\sigma}_i - \vec{\sigma}_0) (\vec{\sigma}_i - \vec{\sigma}_0)^T \qquad M_{\text{sur}}[\sigma_{\text{true}}] = \mathbf{S} \vec{\sigma}_{\text{true}}$$

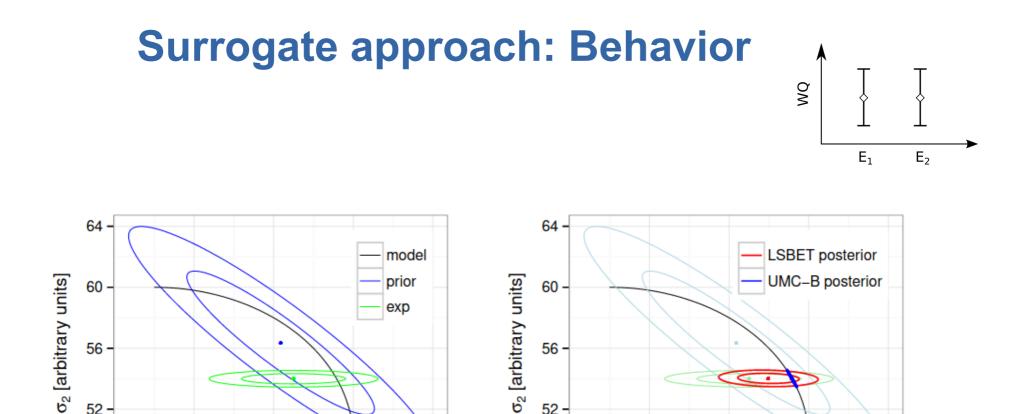


$$\pi(\vec{p}_{\text{true}} | \vec{\sigma}_{\text{exp}}, M_{\text{lin}}) \propto \exp\left[-\frac{1}{2}(\vec{\sigma}_{\text{exp}} - M_{\text{sur}}[\vec{\sigma}_{\text{true}}])^T \mathbf{B}^{-1}(\vec{\sigma}_{\text{exp}} - M_{\text{sur}}[\vec{\sigma}_{\text{true}}])\right] \\ \times \exp\left[-\frac{1}{2}(\vec{\sigma}_{\text{true}} - \vec{\sigma}_0)^T \mathbf{A}_0^{-1}(\vec{\sigma}_{\text{true}} - \vec{\sigma}_0)\right]$$

Analytic update formulas

$$\mathbf{A}_1 = \mathbf{A}_0 - \mathbf{A}_0 \mathbf{S}^T (\mathbf{S} \mathbf{A}_0 \mathbf{S} + \mathbf{B})^{-1} \mathbf{S} \mathbf{A}_0$$

$$\vec{\sigma}_1 = \vec{\sigma}_0 + \mathbf{A}_0 \mathbf{S}^T (\mathbf{S} \mathbf{A}_0 \mathbf{S} + \mathbf{B})^{-1} (\vec{\sigma}_{exp} - M_{sur} [\vec{\sigma}_0])$$



52 -

48 -| 48

52

64

60

 σ_1 [arbitrary units]

56

52

 σ_1 [arbitrary units]

56

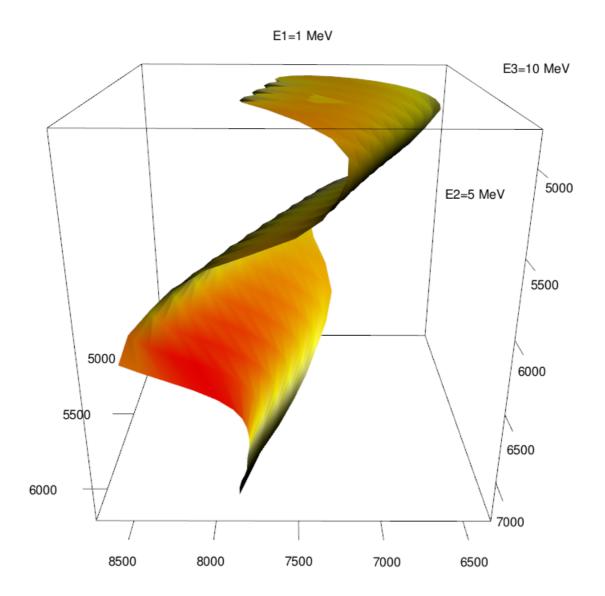
60

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Visualization of non-linearity



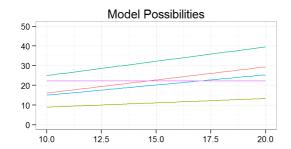
Generalized Least Squares (GLS) in a nutshell

Prior
$$\pi(\vec{p}) = \mathcal{N}(\vec{p} \mid \vec{p_0}, \mathbf{A}_0)$$
Likelihood $f(\vec{\sigma}_{exp} \mid \vec{p}) = \mathcal{N}(\vec{\sigma}_{exp} \mid M_{lin}[\vec{p}], \mathbf{B})$ Bayesian update $\vec{\pi}(\vec{p} \mid \vec{\sigma}_{exp}) \propto f(\vec{\sigma}_{exp} \mid \vec{p}) \pi(\vec{p} \mid \vec{\sigma}_{exp})$ Posterior $\pi(\vec{p} \mid \vec{\sigma}_{exp}) = \mathcal{N}(\vec{p} \mid \vec{p_1}, \mathbf{A}_1)$

$$\vec{p}_1 = \vec{p}_0 + \mathbf{A}_0 \mathbf{S}^T (\mathbf{S} \mathbf{A}_0 \mathbf{S} + \mathbf{B})^{-1} (\vec{\sigma}_{\exp} - M_{\lim} [\vec{p}_0])$$
$$\mathbf{A}_1 = \mathbf{A}_0 - \mathbf{A}_0 \mathbf{S}^T (\mathbf{S} \mathbf{A}_0 \mathbf{S} + \mathbf{B})^{-1} \mathbf{S} \mathbf{A}_0$$

"GLS formulas"

Simple example: straight line model



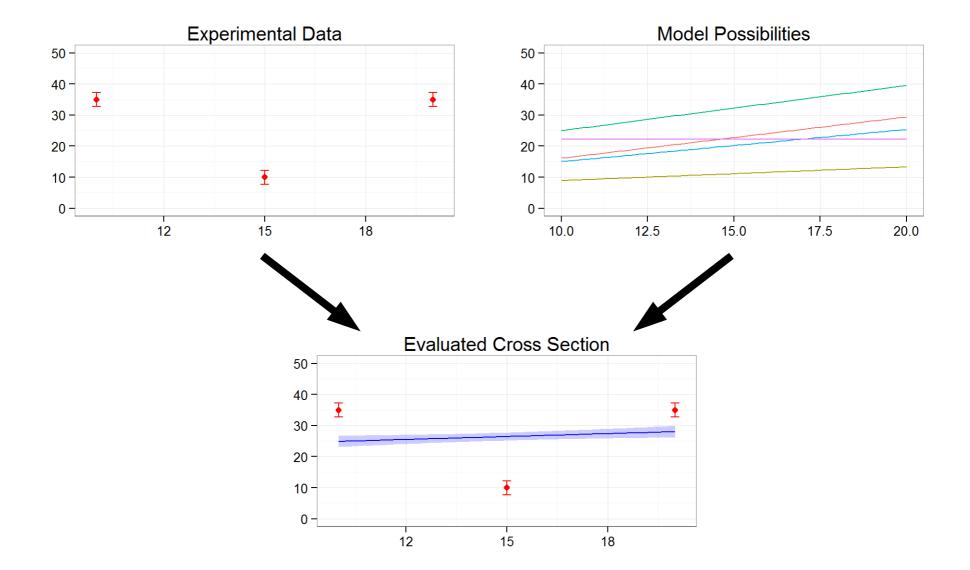
$$y(x) = \mathbf{k}x + \mathbf{d}$$

Prior
$$\frac{k \propto \mathcal{N}(0, \delta_k^2)}{d \propto \mathcal{N}(0, \delta_d^2)} \quad \vec{p}_0 = \begin{pmatrix} 0 \\ 0 \end{pmatrix} \quad \mathbf{A}_0 = \begin{pmatrix} \delta_k^2 & 0 \\ 0 & \delta_d^2 \end{pmatrix}$$

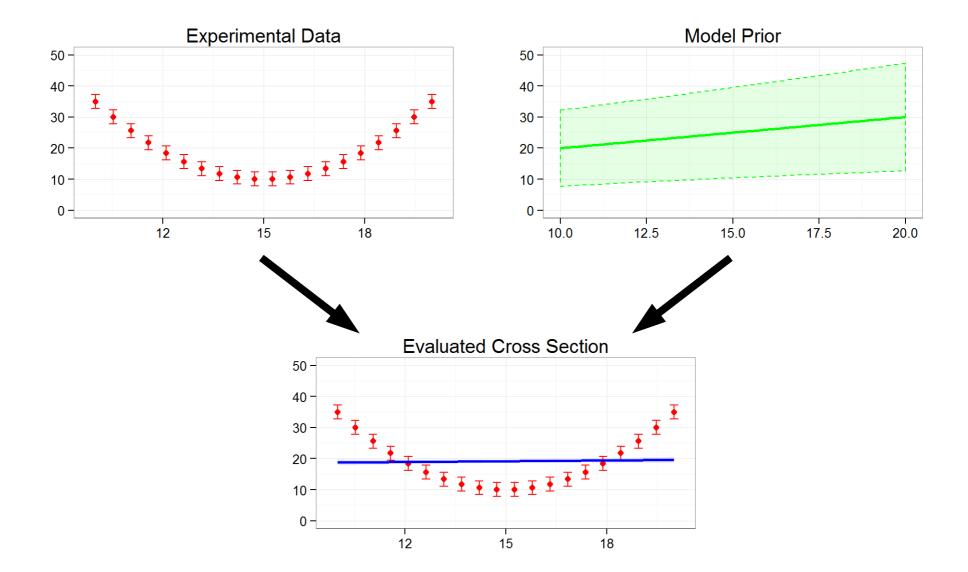
Likelihood
$$\vec{\sigma}_{\exp} = \begin{pmatrix} \sigma_{\exp,1} \\ \sigma_{\exp,2} \\ \vdots \end{pmatrix} \mathbf{B} = \begin{pmatrix} \varepsilon_1 & 0 & 0 \\ 0 & \varepsilon_2 & 0 \\ 0 & 0 & \ddots \end{pmatrix} \overset{\text{Experimental Data}}{\underbrace{}_{\frac{1}{2} & \frac{1}{2} & \frac{1}{3}}}$$

$$M_{\rm lin}[\vec{p}] = \mathbf{S}\vec{p} \qquad \mathbf{S} = \begin{pmatrix} E_1 & 1\\ E_2 & 1\\ \vdots & \end{pmatrix}$$

Straight line model



Straight line model (garbage in, garbage out)



Important insight about models

"All models are wrong but some are useful"

- George E. P. Box





Ice cream or pizza?





Reality



More realistic modeling

Experimental Observation = Model Prediction + Measurement Error

$$\sigma_{Exp}(E) = \sigma_{Mod}(E) + \varepsilon_{Exp}(E)$$

$$\sigma_{True} = \sigma_{Mod}$$

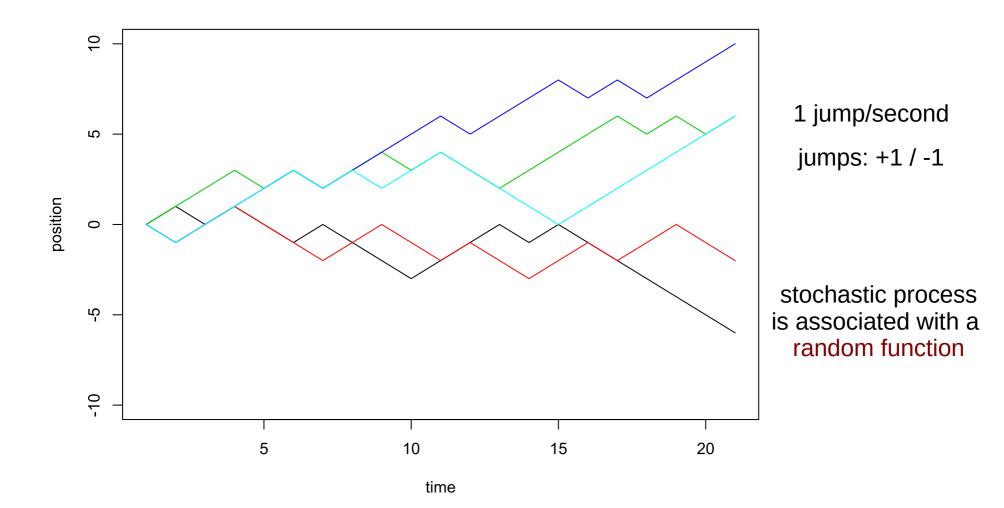
Experimental Observation = Model Prediction + Model Error + Measurement Error

$$\sigma_{Exp}(E) = \sigma_{Mod}(E) + \varepsilon_{Mod}(E) + \varepsilon_{Exp}(E)$$

$$\sigma_{True} = \sigma_{Mod} + \varepsilon_{Mod}$$

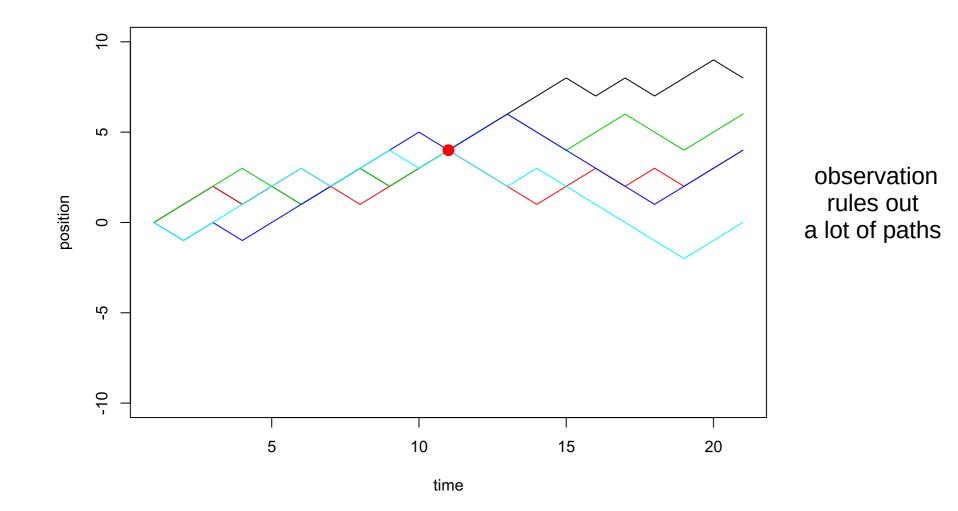
Stochastic process e.g., stock price evolution





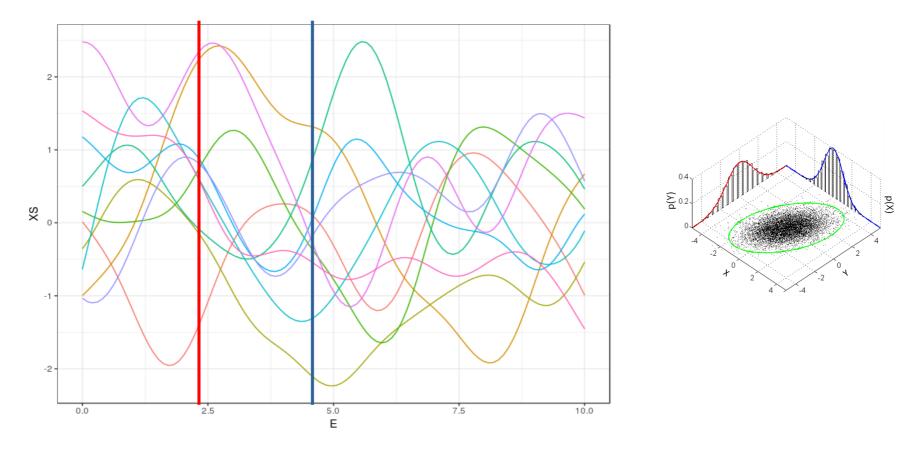
How can we mathematically define a process?

Stochastic process with observation



How can we constrain a process with observations?

Characterization of a Gaussian process



All finite sets of function values at different locations are distributed according to a multivariate normal distribution \rightarrow <u>Gaussian process</u>

Specification of a Gaussian process

Specify a function that yields the covariance between function values $f(\mathbf{x}_1)$ and $f(\mathbf{x}_2)$ for any possible pair \mathbf{x}_1 and \mathbf{x}_2 . This function is called a **covariance function**.

$$\kappa(oldsymbol{x}_1,oldsymbol{x}_2) = \delta^2 \exp\left(-rac{(oldsymbol{x}_1-oldsymbol{x}_2)^2}{2\lambda^2}
ight)$$

Specify another function $\mu(\mathbf{x})$ that yields the center value of the process at location \mathbf{x} . This function is called a **mean function**.

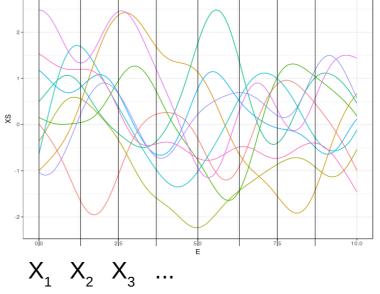
$$\mu(\mathbf{x}) = 0$$

Sampling from a Gaussian process

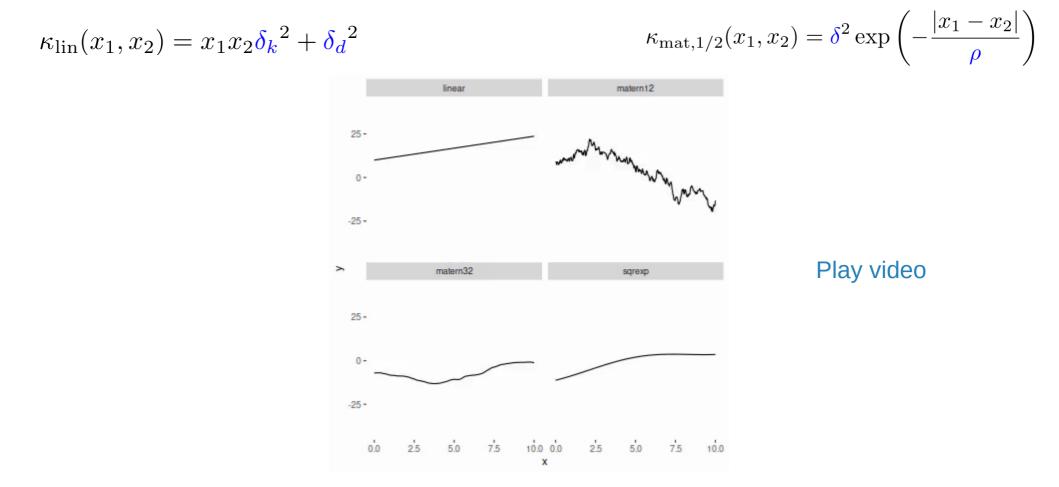
$$\kappa(\boldsymbol{x}_{1},\boldsymbol{x}_{2}) = \delta^{2} \exp\left(-\frac{(\boldsymbol{x}_{1} - \boldsymbol{x}_{2})^{2}}{2\lambda^{2}}\right) \qquad \qquad \mu(\mathbf{x}) = 0$$

$$\mathbf{K} = \begin{pmatrix} \kappa(\boldsymbol{x}_{1},\boldsymbol{x}_{1}) & \kappa(\boldsymbol{x}_{1},\boldsymbol{x}_{2}) & \kappa(\boldsymbol{x}_{1},\boldsymbol{x}_{3}) & \dots \\ \kappa(\boldsymbol{x}_{2},\boldsymbol{x}_{1}) & \kappa(\boldsymbol{x}_{2},\boldsymbol{x}_{2}) & \kappa(\boldsymbol{x}_{2},\boldsymbol{x}_{3}) & \dots \\ \kappa(\boldsymbol{x}_{3},\boldsymbol{x}_{1}) & \kappa(\boldsymbol{x}_{3},\boldsymbol{x}_{2}) & \kappa(\boldsymbol{x}_{3},\boldsymbol{x}_{3}) & \dots \\ \vdots & \vdots & \vdots & \ddots \end{pmatrix} \qquad \vec{m} = \begin{pmatrix} \mu(\boldsymbol{x}_{1}) \\ \mu(\boldsymbol{x}_{2}) \\ \mu(\boldsymbol{x}_{3}) \\ \vdots \end{pmatrix}$$

 $\mathbf{y} \sim \mathcal{N}(\vec{m}, \mathbf{K})$



Examples of Gaussian processes



$$\kappa_{\text{mat},3/2}(x_1, x_2) = \delta^2 \left(1 + \frac{\sqrt{3}|x_1 - x_2|}{\rho} \right) \exp\left(-\frac{\sqrt{3}|x_1 - x_2|}{\rho}\right) \quad \kappa_{\text{sqrexp}}(x_1, x_2) = \delta^2 \exp\left(-\frac{1}{2\lambda^2}(x_1 - x_2)^2\right)$$

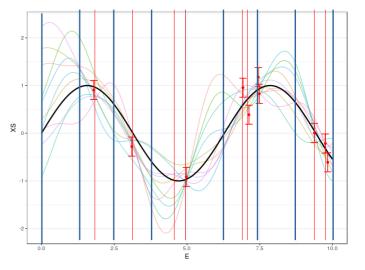
Constraining with observations / Model training

Use GLS formulas:

$$\vec{p}_1 = \vec{p}_0 + \mathbf{A}_0 \mathbf{S}^T (\mathbf{S} \mathbf{A}_0 \mathbf{S} + \mathbf{B})^{-1} (\vec{\sigma}_{\exp} - M_{\lim} [\vec{p}_0])$$
$$\mathbf{A}_1 = \mathbf{A}_0 - \mathbf{A}_0 \mathbf{S}^T (\mathbf{S} \mathbf{A}_0 \mathbf{S} + \mathbf{B})^{-1} \mathbf{S} \mathbf{A}_0$$

Introduce computational mesh ${\bf U}$ and experimental mesh ${\bf V}$ and identify:

$$\begin{array}{ll} \mathbf{A}_{0} \rightarrow \mathbf{K}_{UU} & M_{\mathrm{lin}}[\vec{p}_{0}] \rightarrow \vec{m}_{V} \\ \vec{p}_{0} \rightarrow \vec{m}_{U} & \mathbf{S}\mathbf{A}_{0}\mathbf{S}^{T} \rightarrow \mathbf{K}_{VV} \\ & \mathbf{S}\mathbf{A}_{0} \rightarrow \mathbf{K}_{VU} \end{array}$$

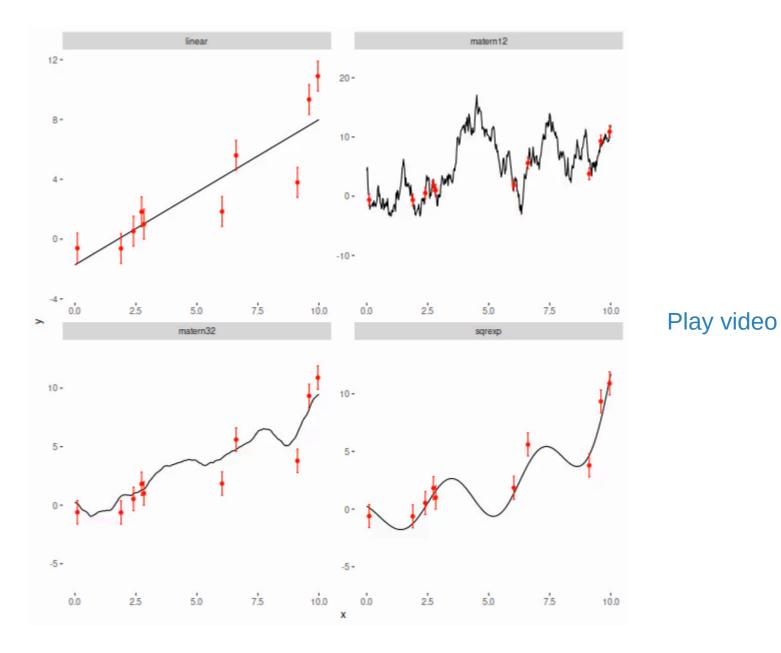


Resulting vector p_1 contains posterior predictions on mesh U and A_1 is the associated covariance matrix

$$\vec{p} \sim \mathcal{N}(\vec{p}_1, \mathbf{A}_1)$$

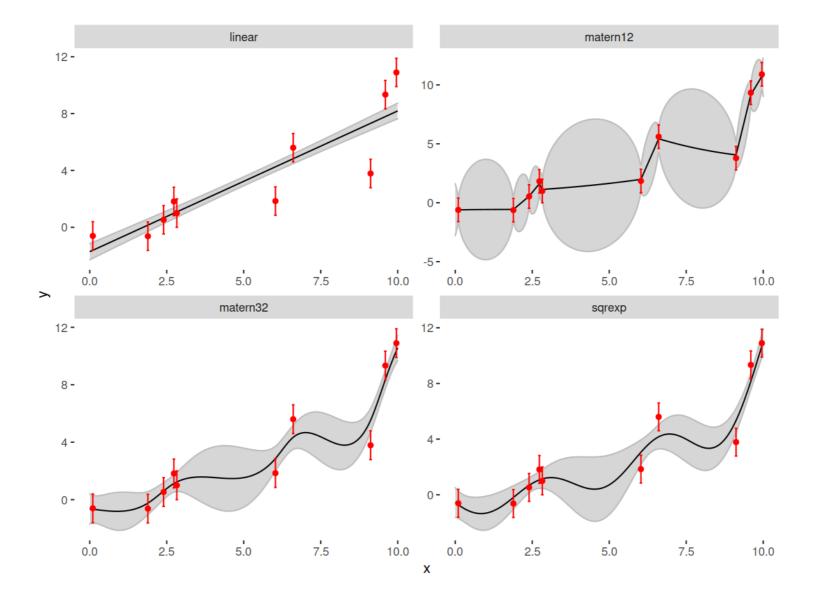
(Other terms: Gaussian process regression, Kriging)

GPs constrained with observations



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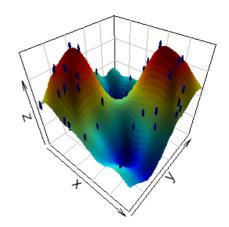
GPs constrained with observations



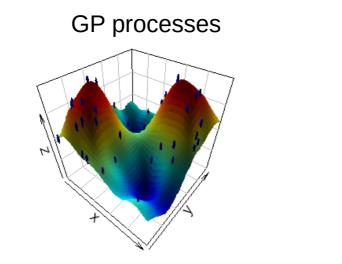
The posterior covariance matrix encodes more than just uncertainties!

Machine learning view on GPs

- Gaussian process regression is a method to learn a non-linear function f(x) from observations {(x_i, y_i)}_{i=1..N}
- It can be used for regression and classification problems
- It is an alternative to other ML techniques, e.g., random forests and neural networks with its own pros & cons



Comparison to neural networks



artificial neural networks

Both approaches ...

- ... are methods for classification and regression
- ... are universal function approximators

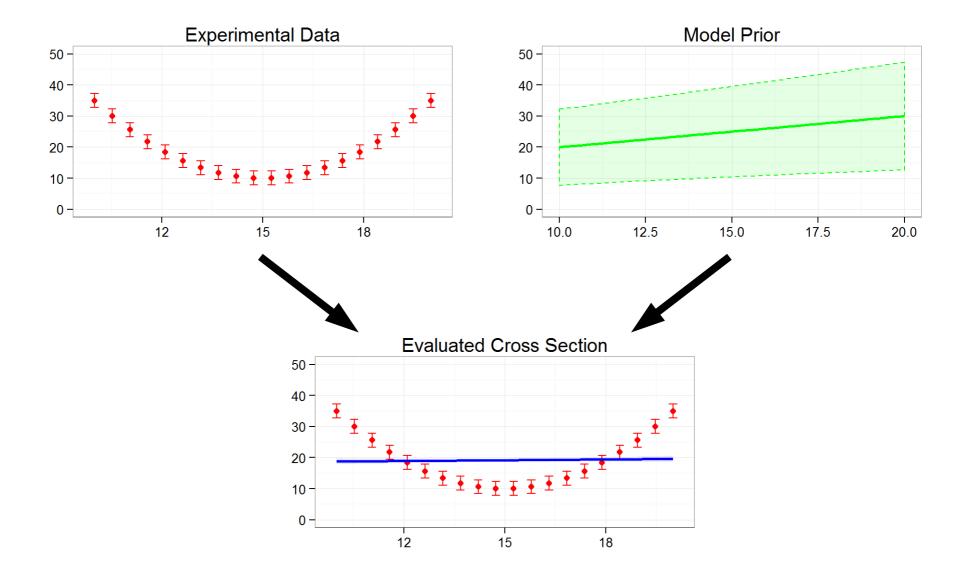
Neural networks ...

- ... scale better to large data sets
- ... are able to capture non-local features
- ... are difficult to interpret

<u>GP processes</u> ...

- ... are statistical methods from the ground up (uncertainties)
- ... facilitate the incorporation of prior assumptions
- ... interface well with existing nuclear data evaluation methods

Why we started talking about GPs



Linear model as Gaussian process

$$y(x) = kx + d$$
 $\vec{p_0} = \begin{pmatrix} 0 \\ 0 \end{pmatrix}$ $\mathbf{A}_0 = \begin{pmatrix} \delta_k^2 & 0 \\ 0 & \delta_d^2 \end{pmatrix}$

$$M_{\text{lin}}[\vec{p}] = \mathbf{S}\vec{p} \qquad \mathbf{S} = \begin{pmatrix} x_1 & 1 \\ x_2 & 1 \\ \vdots & \end{pmatrix} \qquad \mathbf{S}(x) = \begin{pmatrix} x & 1 \end{pmatrix}$$

$$\kappa_{\mathrm{mod}}(x_1, x_2) = \mathbf{S}(x_1) \mathbf{A}_0 \mathbf{S}(x_2)^T$$

Relationship: linear models and GPs

- All linear models (splines, Fourier series, other polynomials, ...) can be represented as a Gaussian process if one assigns a prior covariance matrix to the model parameters/coefficients
- Some Gaussian processes can be represented as a linear model with a finite number of parameters and a multivariate normal prior on the parameters
- All Gaussian processes can be represented as a linear model with an infinite number of parameters

Clarification: e.g., $y(x) = ax^2 + bx$ is a linear model in our context because the predictions at all x are a linear function of the parameters a and b

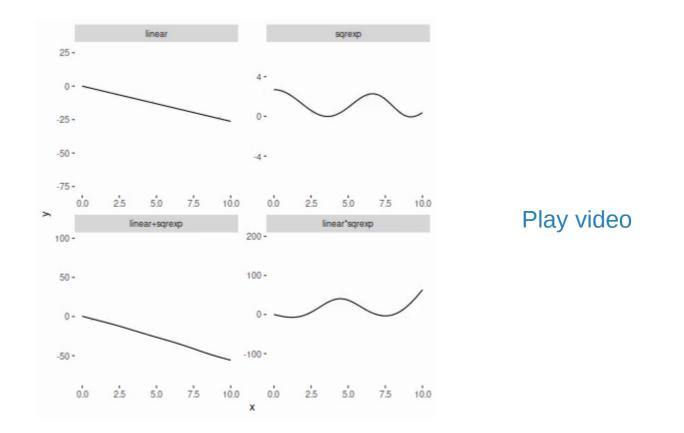
Combination of Gaussian processes

$$\kappa_{1+2}(x_1, x_2) = \kappa_1(x_1, x_2) + \kappa_2(x_1, x_2)$$



$$\kappa_{1\times 2}(x_1, x_2) = \kappa_1(x_1, x_2) \times \kappa_2(x_1, x_2)$$

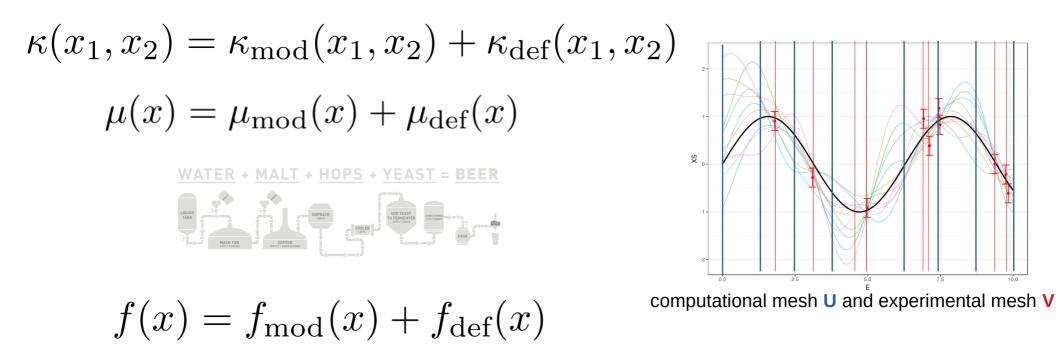
Behavior of combined GPs



$$\kappa_{1+2}(x_1, x_2) = \kappa_1(x_1, x_2) + \kappa_2(x_1, x_2) \implies \text{Sum of random functions } \mathbf{f_1(x) + f_2(x)}$$

$$\kappa_{1\times 2}(x_1, x_2) = \kappa_1(x_1, x_2) \times \kappa_2(x_1, x_2) \implies \text{Some kind of modulation}$$

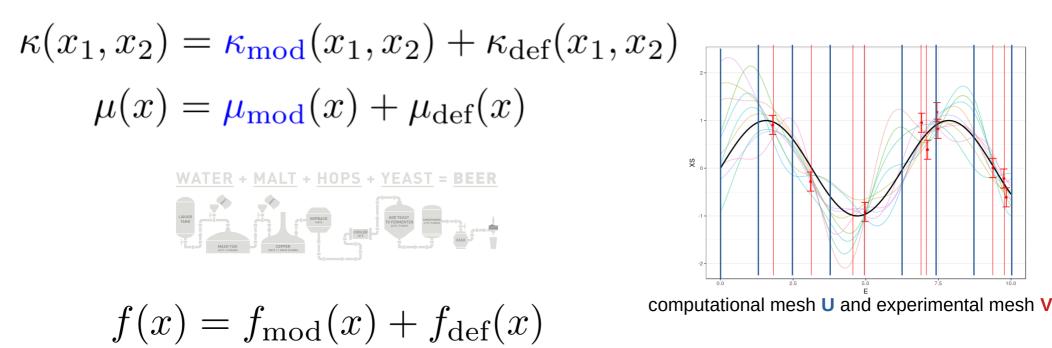
Disentangling Gaussian processes



Compound prediction on computational mesh **U** and associated covariance matrix:

$$\vec{m}'_U = \vec{m}_U + \mathbf{K}_{UV} \left(\mathbf{K}_{VV} + \mathbf{B} \right)^{-1} \left(\vec{\sigma}_{\exp} - \vec{m}_V \right)$$
$$\mathbf{K}'_{UU} = \mathbf{K}_{UU} - \mathbf{K}_{UV} \left(\mathbf{K}_{VV} + \mathbf{B} \right)^{-1} \mathbf{K}_{UV}^T$$

Disentangling Gaussian processes



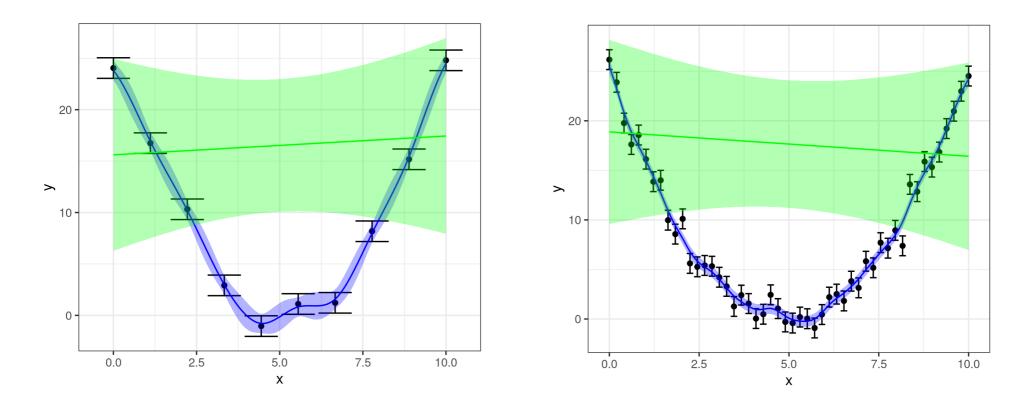
Model prediction on computational mesh U and associated covariance matrix:

$$\vec{m}_{U}^{\text{mod}\prime} = \vec{m}_{U}^{\text{mod}} + \mathbf{K}_{UV}^{\text{mod}} \left(\mathbf{K}_{VV} + \mathbf{B}\right)^{-1} \left(\vec{\sigma}_{\exp} - \vec{m}_{V}\right)$$
$$\mathbf{K}_{UU}^{\text{mod}\prime} = \mathbf{K}_{UU}^{\text{mod}} - \mathbf{K}_{UV}^{\text{mod}} \left(\mathbf{K}_{VV} + \mathbf{B}\right)^{-1} \left(\mathbf{K}_{UV}^{\text{mod}}\right)^{T}$$

Example of disentangled compound GP

$$\kappa_{\text{mod}}(x_1, x_2) := \kappa_{\text{lin}}(x_1, x_2) = x_1 x_2 \delta_k^2 + \delta_d^2$$
$$\kappa_{\text{def}}(x_1, x_2) := \kappa_{\text{sqrexp}}(x_1, x_2) = \delta^2 \exp\left(-\frac{1}{2\lambda^2}(x_1 - x_2)^2\right)$$

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More conservative model prediction uncertainties

Recap

- Introduction to Bayesian statistics with multivariate normal model
- Common approaches in Nuclear Data to (approximately) evaluate the Bayesian update equation (GLS method)
- Gaussian processes as flexible modeling framework to fit functions which is mathematically compatible with the GLS method (Gaussian process regression = GLS with linear models with possibly infinitely many parameters)

Bayesian inference is a large research field and it is applied in many domains, e.g., natural language processing, image analysis, time series prediction, etc. and there is a vast world of possibilities beyond the GLS method, also for nuclear data!

Nuclear data challenges & perspectives

Regarding nuclear models

- Non-linear (nuclear physics is complicated)
- Not analytic (differential equations, simulation)
- Computationally expensive (minutes to hours)

Regarding linear algebra

• Computing with large covariance matrices

Regarding statistical models

- Imperfect physics model
- Multivariate normal distribution may be not always appropriate
- Uncertainties wrong or unknown

Machine Learning techniques

- Robust outlier detection
- Global approaches to cross section predictions over the nuclide chart
- Enhance traditional Bayesian evaluation techniques with ML methods

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References

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